

Accounting for Linkage Disequilibrium in Association Analysis of Diverse Populations

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ABSTRACT: The National Human Genome Research Institute's catalog of published genome-wide association studies (GWAS) lists over 10,000 genetic variants collectively associated with over 800 human diseases or traits. Most of these GWAS have been conducted in European-ancestry populations. Findings gleaned from these studies have led to identification of disease-associated loci and biologic pathways involved in disease etiology. In multiple instances, these genomic findings have led to the development of novel medical therapies or evidence for prescribing a given drug as the appropriate treatment for a given individual beyond phenotypic appearances or socially defined constructs of race or ethnicity. Such findings have implications for populations throughout the globe and GWAS are increasingly being conducted in more diverse populations. A major challenge for investigators seeking to follow up genomic findings between diverse populations is discordant patterns of linkage disequilibrium (LD). We provide an overview of common measures of LD and opportunities for their use in novel methods designed to address challenges associated with following up GWAS conducted in European-ancestry populations in African-ancestry populations or, more generally, between populations with discordant LD patterns. We detail the strengths and weaknesses associated with different approaches. We also describe application of these strategies in follow-up studies of populations with concordant LD patterns (replication) or discordant LD patterns (transferability) as well as fine-mapping studies. We review application of these methods to a variety of traits and diseases.

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Introduction

Linkage disequilibrium (LD) refers to the tendency of alleles at linked loci to be inherited together [Pritchard and Przeworski, 2001]. In evolutionary terms, African-ancestry populations are the oldest of modern human populations. Compared to European- or Asian-ancestry populations, African-ancestry populations have experienced more generations for LD to decay [Tishkoff et al., 1996]. Furthermore, bottlenecks associated with Out-of-Africa migrations have led to reduced haplotypic diversity in European- and Asian-ancestry populations relative to African-ancestry populations (Fig. 1) [Campbell and Tishkoff, 2008]. Consequently, as a function of demography and time, African-ancestry individuals tend to have shorter segments of linked alleles whereas European- or Asian-ancestry individuals tend to have longer segments of linked alleles. The relative lengths of these segments create distinct patterns that can be leveraged in genomic studies.

Genome-wide association studies (GWAS) use tag variants to locate regions in the genome in which phenotype-associated variants reside. This strategy is based on the com-

mon disease-common variant hypothesis, which assumes that tag variants associated with a phenotype are in LD with causal variants [Cantor et al., 2010]. European- and Asian-ancestry populations with strong LD are ideal for conducting discovery analyses whereas African-ancestry populations with weak LD are better suited for fine-mapping [Gabriel et al., 2002]. African-ancestry populations are also well suited for discovery of novel loci, due to differences in LD structure and/or allele frequencies. Several studies provide examples in which associated variants were discovered in African-ancestry populations [Bensen et al., 2012; Dumitrescu et al., 2011; Freedman et al., 2010]. The application of LD to derive information regarding genetic epidemiology is not surprising given the fact that LD provides insight into the history of populations, including their mating patterns, geographic subpopulation structure, natural selection, gene conversion, mutation, and changes in allele frequencies over time [Weir, 2008].

GWAS are a valuable tool for establishing correlations between genetic variants and phenotypes, with human disease- or trait-associated single nucleotide polymorphisms (SNPs) exceeding 10,000 [Hindorf et al., 2009]. GWAS have led to identification of disease-associated loci and biologic pathways involved in disease development [Haines et al., 2005; Klein et al., 2005]. Some of these findings have led to the development of novel medical therapies [Ashtari et al., 2011; Hayes et al., 2007; Los et al., 2007] or evidence for

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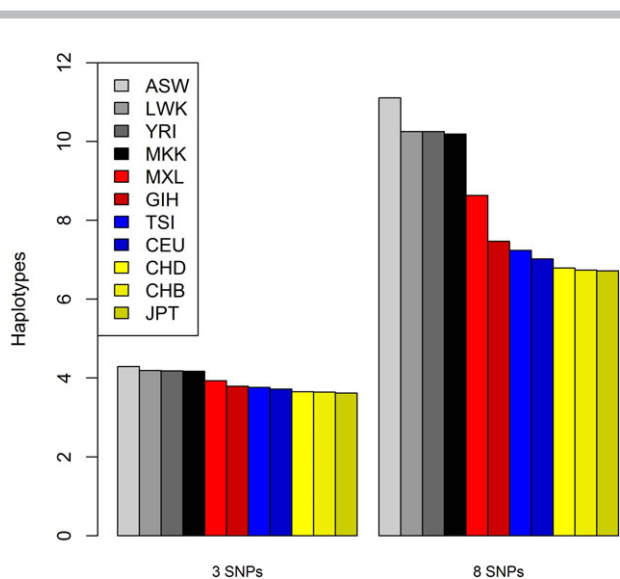


Figure 1. Global haplotype diversity. The maximum number of haplotypes for a window of three consecutive SNPs is eight and for eight SNPs is 256. All counted haplotypes have an estimated frequency $\geq 1\%$ in the HapMap Phase III data. Population descriptors: ASW = African ancestry in Southwest USA; LWK = Luhya in Webuye, Kenya; YRI = Yoruba in Ibadan, Nigeria; MKK = Maasai in Kinyawa, Kenya; MXL = Mexican ancestry in Los Angeles, California; GIH = Gujarati Indians in Houston, Texas; TSI = Tuscan in Italy; CEU = Utah residents with Northern and Western European ancestry from the CEPH collection; CHD = Chinese in metropolitan Denver, Colorado; CHB = Han Chinese in Beijing, China; JPT = Japanese in Tokyo, Japan. As expected, African ancestry populations have more haplotype diversity.

prescribing a particular drug as the appropriate treatment for a given patient beyond phenotypic appearances or socially defined constructs of race or ethnicity [Rotimi and Jorde, 2010]. Such findings will have an impact on the treatment of disease in populations around the world. Although most GWAS have been conducted in European-ancestry populations [Need and Goldstein, 2009], more diverse populations are increasingly being included in GWAS [Adeyemo et al., 2009; Charles et al., 2011; Norris et al., 2009; Rich et al., 2009; Tin et al., 2011].

Coverage

The tag SNPs on a genome-wide association array have varying amounts of coverage based on the ancestral origins of the population upon which the array's development was based and the population upon which it is being used. LD is central in identifying the tag variants representing the genomes of various populations. Early SNP arrays were largely based on SNP discovery in individuals of European ancestry with the recognition that most of the tag variants would be portable to Asian-ancestry populations but that African-ancestry populations would require supplementation with additional markers [Conrad et al., 2006; Need and Goldstein, 2006]. Coverage of common variation (minor allele frequency $\geq 5\%$) in the HapMap Phase II data by these early GWAS panels at $r^2 \geq 0.08$ ranged from 75–86%

for CEU samples to 28–49% for YRI samples [Barrett and Cardon, 2006]. Newer arrays including data from all three phases of HapMap as well as variants identified by the 1000 Genomes Project achieve coverage of SNPs with a minor allele frequency of $\geq 1\%$ of 93%, 92%, and 76% for the CEU, CHB+JPT, and YRI populations, respectively [Delano et al., 2010]. Population-specific arrays such as the Axiom Genome-Wide PanAFR Array Plate Set achieve coverage of $\sim 90\%$ for common variants and $\sim 85\%$ for low-frequency variants (minor allele frequency $\geq 2\%$) for the African-ancestry HapMap YRI, LWK, MKK, and ASW samples [Affymetrix, 2011].

Tag variants on SNP arrays are chosen to reduce redundancy as measured by r^2 . Genotypes for markers not on an array can be estimated using imputation, which is essentially the inverse of tagging. Imputation fills in missing genotypes by relying on a reference set of haplotypes; the weaker LD is, the less confidently missing genotypes can be imputed [Marchini and Howie, 2010]. Consequently, imputation performs worse in African-ancestry populations compared to European- or Asian-ancestry populations [Huang et al., 2005].

Commonly used measures of LD have been thoroughly reviewed in prior publications [Slatkin, 2008; Weir 2008]; however it is important that this content be briefly reviewed as groundwork for our discussion regarding uses of LD to interrogate the genome between diverse populations and its application to GWAS, GWAS follow-up, fine-mapping, and rare variant analysis. A brief review of measures of LD is provided in the Supporting Information Appendix A.

Challenges With Using LD to Interrogate the Genome

A replicate sample is defined as a random sample from the same population as the discovery sample. This definition implies that LD structure is concordant between samples. If a follow-up sample has discordant LD structure compared to the discovery sample, the definition of replication is not met and the term *transferability* is used [Shriner et al., 2009]. When transferability fails, the variant originally associated with the trait in the discovery population may be population specific [Shriner et al., 2009].

Many follow-up studies have resulted in failure to find association for the exact SNP identified in the discovery population. The phrase exact replication refers to this strategy [Clarke et al., 2007; Colhoun et al., 2003; Ioannidis et al., 2001; Ioannidis, 2007]. Some follow-up attempts have led to identification of different loci being associated with the disease or trait while others have led to the identification of different diseases or traits being associated with the loci in different populations or differences in the direction of the effect of the associated loci [Clarke et al., 2007; Ober, 2005]. These instances may be due in part to discordant LD structure between populations [Clarke et al., 2007; Ong and Teo, 2010] and allelic, genetic, or trait heterogeneity [Clarke et al., 2007]. As investigators began to examine findings from GWAS from diverse populations, discordant LD structure between populations would often yield false negative findings, leaving

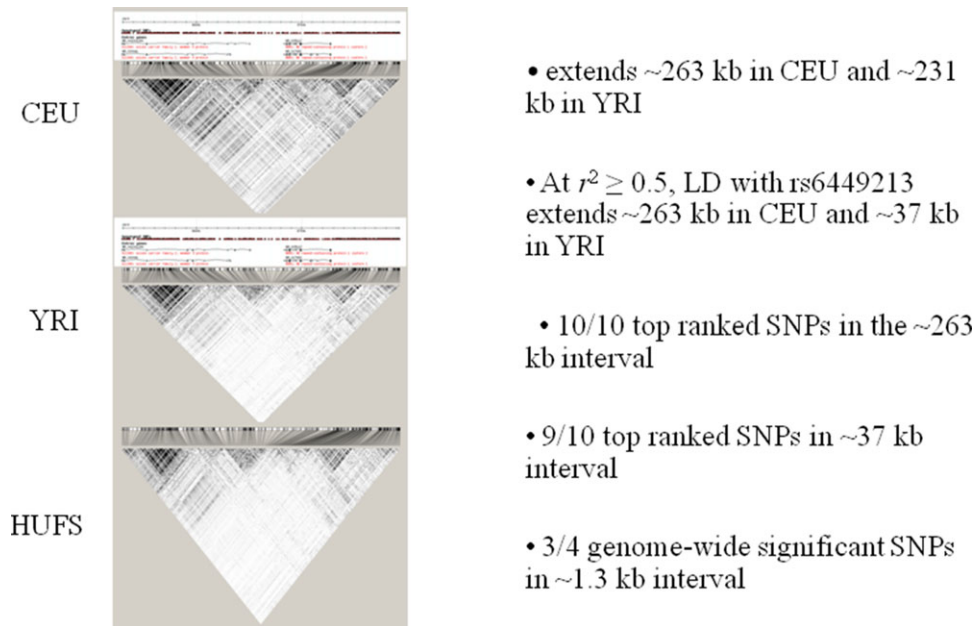


Figure 3. Comparison of linkage disequilibrium in the HapMap Phase II CEU, YRI, and HUFS sample (African Americans). Using SNPs with $r^2 \geq 0.5$, the LD interval encompassing the index SNP identified in the European-ancestry discovery population is larger in the HapMap CEU sample than the LD interval encompassing the index SNP in the HapMap YRI sample. Taking advantage of this discordant pattern of LD results in a reduction from 263 kb in CEU to 37 kb in YRI to 1.3 kb for 75% of the genome-wide significant SNPs in the African American sample, illustrating fine-mapping when LD is weaker in the follow-up population [Ashtari et al., 2011].

et al., 2008] was the only gene identified in the African American sample. The association of the index SNP rs6449123 in *SLC2A9* with serum uric acid was fine mapped in the African American, CEU and YRI samples [Charles et al., 2011]. The LD pattern for the African American sample was almost identical to that of the YRI sample. At a threshold of $r^2 \geq 0.5$ for SNPs in LD with rs6449123, the associated interval was 263 kb in the CEU sample while in the African American sample the associated interval was 37 kb (Fig. 3) [Charles et al., 2011]. Three of four SNPs associated with genome-wide significance in the African American sample lay within a 1.3 kb interval [Charles et al., 2011]. These results provide compelling evidence of the advantage of following up GWAS signals originally identified in populations with strong LD (European ancestry) in populations with weak LD (African ancestry). Similar methodologies have been used for fasting plasma glucose, coronary artery disease, and type 2 diabetes and are described below.

Fasting Plasma Glucose

Exact and local methods were applied to a study of fasting plasma glucose in African Americans [Ramos et al., 2011]. Replication was assessed for 20 SNPs and an additional 9 SNPs were assessed using local methods [Ramos et al., 2011]. The region of assessment for local transferability was determined using $r^2 \geq 0.3$ for SNPs within 250 kb of the index SNPs [Ramos et al., 2011]. SNPs were selected using the HapMap LD data from the CEU population with the exception of

rs2166706, for which the reference sample was GIH to account for the discovery of this association in Asian Indians [Ramos et al., 2011]. A sample covariance matrix based on genotypes was decomposed and the effective degrees of freedom were denoted as $N_{eff} = \frac{(\sum_{k=1}^K \lambda_k)^2}{\sum_{k=1}^K \lambda_k^2}$, in which λ_k is the k^{th} eigenvalue from the $k \times k$ covariance matrix for the K SNPs in the region [Bretherton et al., 1999]. The threshold for significance was defined as $\alpha = \frac{0.05}{N_{eff}}$, corresponding to a partial Bonferroni correction rather than the full Bonferroni correction of $\alpha = \frac{0.05}{K}$ since $1 \leq N_{eff} \leq K$ [Ramos et al., 2011]. Using this methodology, associations for SNPs in or near five genes were replicated and SNPs in four additional genes were found to be associated using transferability analysis [Ramos et al., 2011].

Coronary Heart Disease

Lettre et al. [2011] fine mapped in African Americans five loci discovered in European-ancestry individuals as associated with coronary artery disease and its risk factors. SNPs correlated ($r^2 \geq 0.5$) to the index SNP in the HapMap CEU sample but uncorrelated ($r^2 \leq 0.1$) in the HapMap YRI sample were evaluated [Lettre et al., 2011]. For each of these five loci, markers more strongly associated with the phenotype than the index SNPs were discovered in the African American sample [Lettre et al., 2011]. As an example, the index SNP rs174547 from a European-ancestry population is in strong LD with rs1535, the most strongly associated SNP in the locus

among African Americans. The two markers are in strong LD in the HapMap CEU sample ($r^2 = 1$) but not in the HapMap YRI sample ($r^2 \geq 0.09$) [Lettre et al., 2011]. LD surrounding rs174547 extends over 113 kb in the HapMap CEU sample but rs1535 is not strongly correlated with any other marker in the HapMap YRI sample [Lettre et al., 2011].

Type 2 Diabetes

Palmer et al. [2012] evaluated 37 index SNPs associated with type 2 diabetes in African Americans. Sets of SNPs surrounding the index SNPs were based on LD blocks defined by D' using the HapMap CEU data [Gabriel et al., 2002]. Locus-specific association P -values were adjusted for the effective number of SNPs as calculated using the method of Li and Ji [2005]. One index SNP was significantly associated using the exact approach and one additional SNP was significantly associated using a local approach [Palmer et al., 2012].

A worked example for serum bilirubin [Chen et al., 2011; Johnson et al., 2009; Kang et al., 2010; Sanna et al., 2009] is provided in Appendix 2 Text Box.

Contrasting LD in Two Samples

If correlations between a pair of SNPs in two populations result in D' values that are positive in one sample and negative in another sample, both the type I and type II error rates in association studies can increase [Teo et al., 2009b]. To quantify variation in LD between samples, Teo et al. devised a signed r^2 statistic defined as $(-1)^{I(P_{AB} < P_A P_B)} \frac{(P_{AB} - P_A P_B)^2}{P_A P_B P_B P_B}$, in which $I(P_{AB} < P_A P_B)$ is an indicator function equaling one when true and equaling zero otherwise [Ong and Teo, 2010; Teo et al., 2009a; Teo et al., 2009b]. This statistic is the conventional r^2 statistic with the sign of D' . Matrices M_1 and M_2 are constructed using the signed r^2 values between all pairs of SNPs for samples 1 and 2. Let Δ_i represent the diagonal matrix of sorted eigenvalues from the decomposition of M_i . The VarLD score is defined as $\text{trace} |\Delta_1 - \Delta_2|$, with significance of the score assessed by sampling with replacement [Teo et al., 2009a]. Examples of regions for which the VarLD score indicates variation in LD between samples include *VKORC1*, *NRG1*, *LCT*, *SLC24A5*, *HBB*, and *MHC* [Ong and Teo, 2010].

The VarLD score can empirically demonstrate that LD variation in specific genomic regions is associated with variation in haplotype frequencies and that nontransferability may be more common in these regions [Teo et al., 2009a]. Additionally, loci under different selection pressures in two populations are more likely to have discordant LD patterns between populations [Teo et al., 2010]. A weakness of the VarLD score is exclusion of the eigenvectors, which leads to a loss of power. To illustrate, suppose two samples have LD patterns that are mirror images (Fig. 4). The VarLD score is 0 because the sorted eigenvalues are identical. However, if the eigenvectors are first matched by pairing each eigenvector from sample 1 with the most closely parallel eigenvector from sample 2 (based on the angle between two vectors equaling the arccosine of the dot product of two orthonormal eigenvectors), then the VarLD score reaches its maximum value.

Krzanowski used a two-stage approach assessing both eigenvalues and eigenvectors to assess the equality of matrices of the correlation coefficients [Krzanowski, 1979; Krzanowski, 1993]. Zaykin also uses this two-stage approach, while calculation of the VarLD score includes only assessment of the eigenvalues (calculating the absolute difference of the eigenvalues). There are many well-established tests of equality that have been described in textbooks on multivariate statistical analysis [Anderson, 1958], however, we have not observed instances of their application to use of LD to interrogate the genomes of diverse populations during this review. It is not clear why these classical approaches have not been used in the field to date.

Multivariate Normality

Zaitlen et al. [2010] described an approach for testing association for multiple correlated SNPs from diverse populations. For each study, association Z scores are assumed to follow the multivariate normal distribution, $Z \sim \text{MVN}(\lambda\sqrt{n}, \Sigma)$, in which $\lambda\sqrt{n}$ is the vector of noncentrality parameters and Σ is the matrix of haplotype-based correlation coefficients [Zaitlen et al., 2010]. The Z scores are weighted across studies, yielding test statistics \mathbf{T} that are also multivariate normal, $\mathbf{T} \sim \text{MVN}(\frac{\sum \sqrt{n}\lambda\sqrt{n}}{\sqrt{\sum n}}, \frac{\sum n\Sigma}{\sum n})$ [Zaitlen et al., 2010]. Clark et al. [2010] described a Bayesian test comparing the hypothesis that two covariance matrices are equal to the hypothesis that two covariance matrices are different. Due to the assumption of multivariate normality and the use of conjugate priors, the Bayesian test can be computed efficiently in closed form and therefore does not require permutation testing.

Metaanalysis

Teo et al. [2010] compared inverse-variance weighted fixed effects meta-analysis with Fisher's combined P -value to fine map disease-associated variants. Between-study heterogeneity (effect size variation) violates the assumption of uniformity of effect sizes required for a fixed effects model to be most powerful [Cochran, 1954; de Bakker, Ferreira et al., 2008; Mantel and Haenszel, 1959]. For populations with discordant LD structure, a random effects model is more appropriate than either of these approaches. However, the commonly employed random effects model effectively assumes between-study heterogeneity under the null hypothesis and yields conservative P -values compared to the fixed effects model. Inclusion of between-study heterogeneity in the alternative hypothesis, rather than in the null hypothesis, results in greater power for detecting association but is also more sensitive to confounding [Han and Eskin, 2011].

Morris developed a Bayesian partition model in which populations are clustered based on shared ancestry [Morris, 2011]. This model is more powerful than the fixed effects model if there is heterogeneity and more powerful than the random effects model if populations are related. Potential effects of LD are not formally part of this model but are

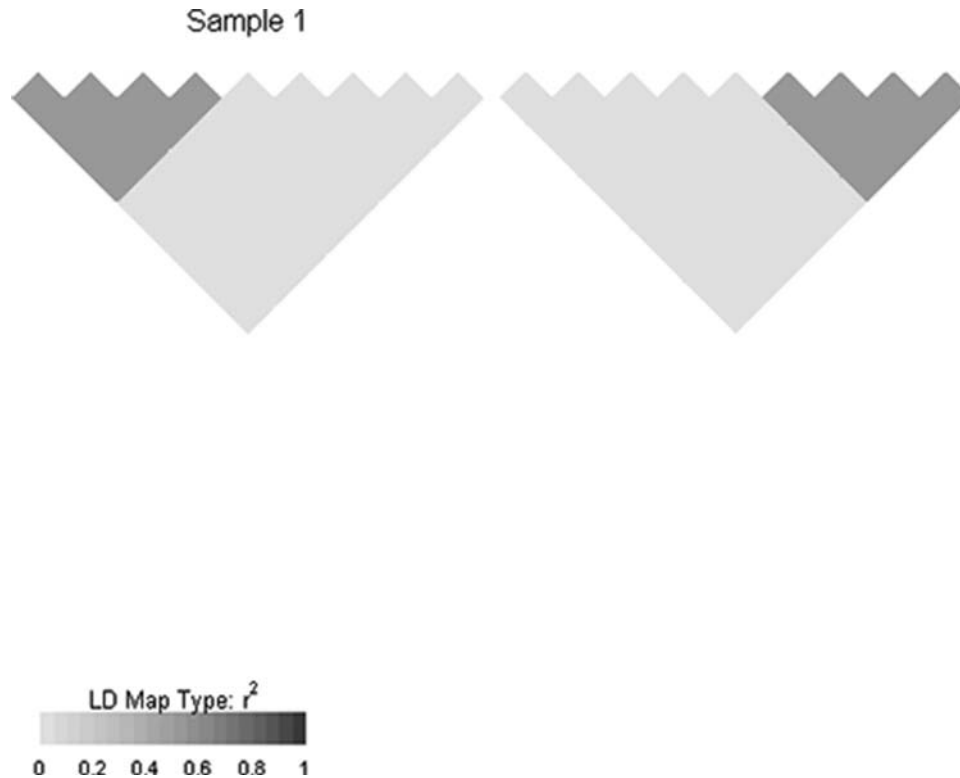


Figure 4. Error associated with using the trace of the eigenvalues without matching the eigenvectors. Sample 1 consists of 10 SNPs with pairwise $r^2 = 0.5$ for all pairs of SNPs 1–5 and $r^2 = 0$ for all other pairs of SNPs. Sample 2 is the mirror image of sample 1, with pairwise $r^2 = 0.5$ for all pairs of SNPs 6–10 and $r^2 = 0$ for all other pairs of SNPs. Without matching the eigenvectors, the sorted eigenvalues are identical.

indirectly captured through cluster-specific allelic effects. Using this approach three of 19 loci associated with type 2 diabetes showed evidence of heterogeneity in allelic effects among European American, African American, Hispanic, Japanese American, and Native Hawaiian ancestry [Morris, 2011].

Rare Variant Analysis

The common disease-common variant hypothesis is the paradigm upon which GWAS were initially designed; however, variants identified by most GWAS account for a small fraction of the heritability of complex diseases under investigation. The common disease-rare variant hypothesis posits that single or multiple rare variants with moderate to large effects are the source of some of the missing heritability. While a review of techniques for rare variant analysis is beyond our scope, two points are relevant. First, a causal rare variant that is in strong LD with a noncausal common variant may result in the (indirect) association of the noncausal common variant with the phenotype [Dickson et al., 2010]. Second, some methods for rare variant analysis assume that rare variants are not in LD whereas other methods relax this assumption, including methods designed to jointly assess common and rare variants [Shriner and Vaughan, 2011; Yuan et al., 2012].

Discussion

In this review, we have provided a comprehensive resource of statistical methods utilizing the most common measures of LD developed for addressing the challenges and opportunities associated with following-up in African-ancestry populations, GWAS initially conducted in European-ancestry populations or, more generically, between populations with discordant LD patterns. We provided examples of the application of these procedures to phenotypes such as coronary heart disease, fasting plasma glucose, serum uric acid, stature, bilirubin, and type 2 diabetes. We also described the strengths and weaknesses of each approach (Table 1). This information provides investigators with a resource for selecting the most appropriate analytic method based on their sample and research question when seeking to compare populations with discordant LD patterns.

Each method adds valuable information to the literature. The LD contrast test was designed to compare pairwise LD matrices between cases and controls while maintaining robustness to departure from HWE. The LD contrast test has been modified for case-parent data, for quantitative traits, and to eliminate confounding caused by background LD. The VarLD score quantifies differences in LD patterns between two samples but requires knowledge of phase and is underpowered. Transferability analysis is an extension of

Table 1. Select methods for leveraging LD in populations with discordant LD patterns

Test	LD contrast test ⁴⁹	Modified LD contrast ⁵¹	VarLD ⁷⁵	Transferability test ⁵³	Transethnic fine-mapping ⁶⁰	Localization success rate ⁷⁷	Bayesian partitioning ⁸²
Study Design	Case/control	Case/control	Case/control	Case/control	Metaanalysis	Metaanalysis	Metaanalysis
Population LD pattern	Concordant	Continuous traits	Continuous traits	Continuous traits	Discordant	Discordant	Concordant/Discordant
Data/Methods	Tag SNP markers, GWAS, sequence	Tag SNP markers, GWAS, sequence	Tag SNP markers, GWAS, sequence	Tag SNP markers, GWAS, sequence	Fine-mapping	Tag SNP Markers	GWAS, Fine-mapping
Strengths	Increased power compared to haplotype-based test or multilocus tests based on SNP scores in populations with concordant LD between cases and controls	Accounts for confounding due to background LD	Models correlation between regional genomic variation and haplotype diversity	Improved variant localization exact, or local replication, transferability and fine mapping	Improved variant localization	Improved variant localization	Robust to departures from allelic heterogeneity of effect
Weaknesses	Robust to departures from HWE	Robust to departures allelic homogeneity of effect	Models correlation between selection pressure and LD	Can be employed for use with multiple samples with different ancestral backgrounds	Can be applied when comparing multiple samples with different ancestral backgrounds	Software available	Models between-population allelic heterogeneity
	Test statistic for comparing LD plots	Less powerful than haplotype-based tests/multilocus tests when the case/control populations have discordant LD	Eigenvectors exclusion resulting in loss of power	Ease of use	Between-study heterogeneity violates assumption of homogeneity resulting in loss of power	Violation of the assumption of multivariate normality results in decreased power	Increased power compared to random and fixed effects models
	Requires permutation testing	NA	Non-GWAS data requires permutation	None			Software available by request
	Confounding due to background LD	Indels, variation in coverage or genotyping quality decrease power					

replication analysis designed to overcome the challenges associated with transferring results of a GWAS from a discovery sample to a follow-up sample with discordant LD. Depending on the samples under investigation, transferability analysis may also be used for fine mapping. This approach often results in remarkable resolution when contrasting genomic findings from European-ancestry populations with findings from African-ancestry populations [Charles et al., 2011; Chen et al., 2011; Ramos et al., 2011]. Random effects models of metaanalyses have been described but have not yet achieved widespread use.

An increasing number of genomics projects, such as the 1000 Genomes Project, the Singapore Genome Variation Project, and the Human Heredity and Health in Africa Initiative (H3Africa.org), include non-European populations. This type of research requires strategies for addressing the analytic challenges spawned from the diversity inherent in human genomes. Resources presented in this review will help to ensure that results are not marred by false positive or false negative findings. The analytic methods incorporating LD described in this review are applicable to the analysis of both genotype and sequence data and should maintain their relevance for the foreseeable future.

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