

Spider mites example

Dose of DDT	No. survived	No. dead
0.0	18	7
0.5	19	6
1.0	12	13
1.5	5	20
2.0	6	19
2.5	2	23
3.0	1	24

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Binary vs. continuous outcomes

Continuous: ANOVA $\longleftrightarrow$ Regression

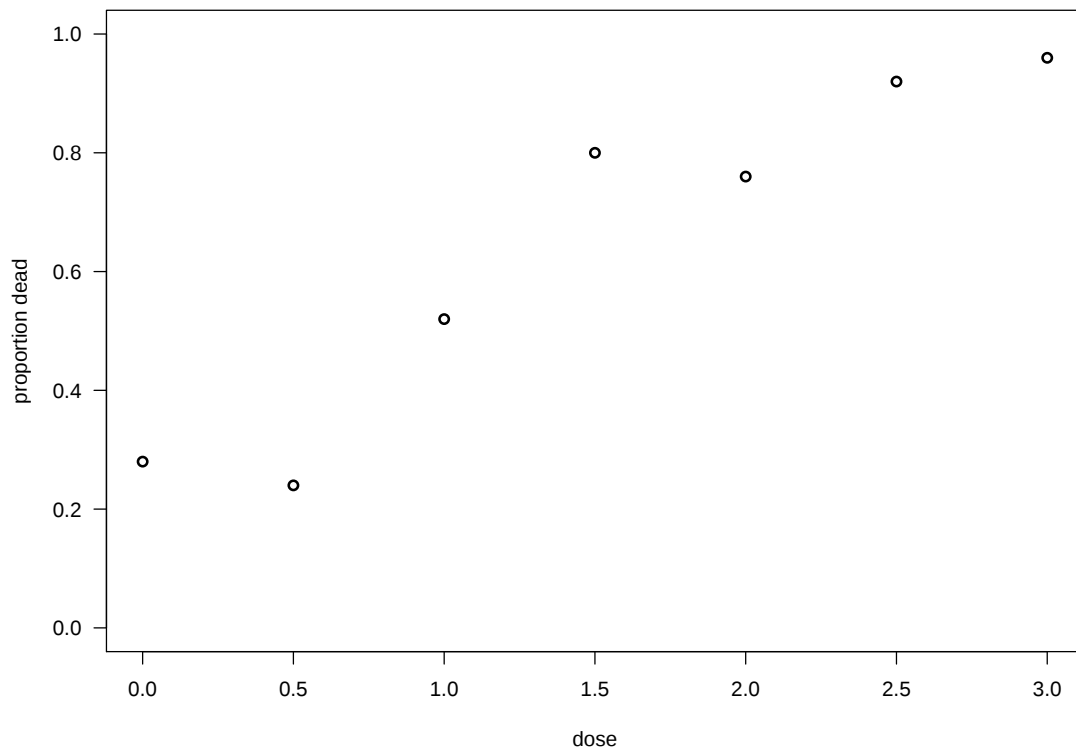
Binary: $k \times 2$ table $\longleftrightarrow$?

Goals:

- Relationship between dose and $\text{Pr}(\text{dead})$.
- Dose at which $\text{Pr}(\text{dead}) = 1/2$.

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A plot of the data



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Linear regression

Model:

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \cdots + \beta_k x_k + \epsilon, \quad \epsilon \sim \text{iid Normal}(0, \sigma^2)$$

This implies:

$$E(y \mid x_1, \dots, x_k) = \beta_0 + \beta_1 x_1 + \cdots + \beta_k x_k$$

β_i = increase in mean of Y associated with a unit change in x_i

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Binary outcomes

Let $p_d = \Pr(\text{dead} \mid \text{dose } d)$

$$p_d = \beta_0 + \beta_1 d \quad ?$$

$$0 \leq p_d \leq 1 \quad \text{but} \quad -\infty \leq \beta_0 + \beta_1 d \leq \infty$$

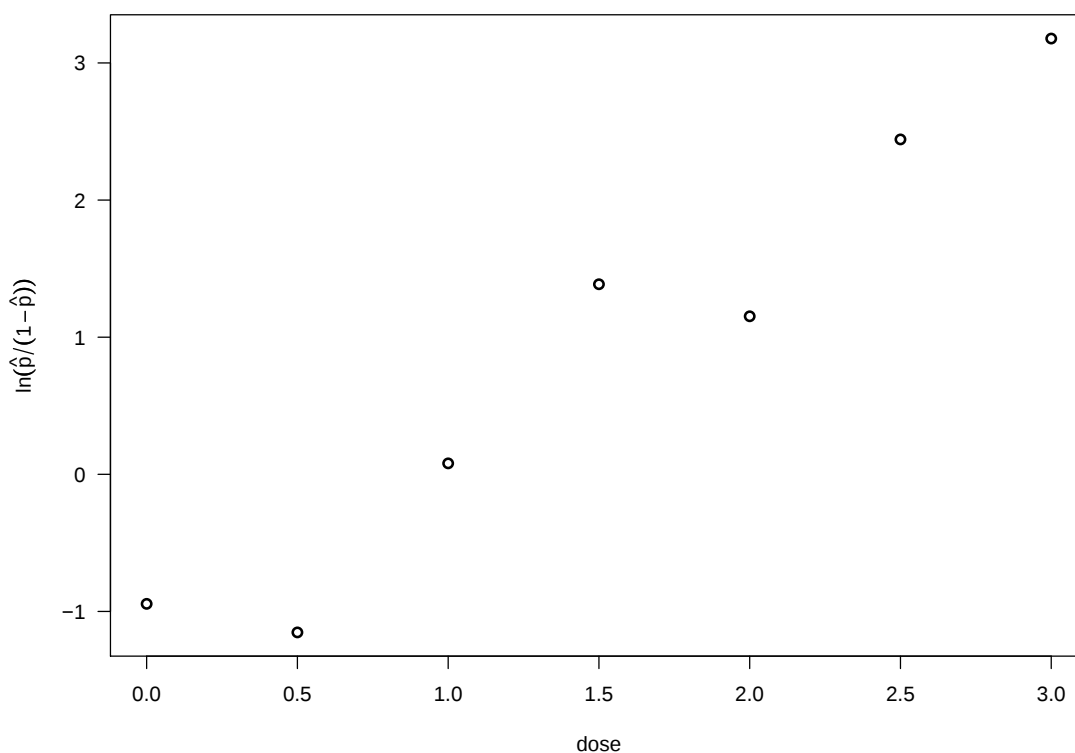
Odds of death: $0 \leq \frac{p_d}{1 - p_d} \leq \infty$

log odds of death: $-\infty \leq \ln \left(\frac{p_d}{1 - p_d} \right) \leq \infty$

$\ln \left(\frac{p}{1 - p} \right)$ is also called **logit(p)** or the **logistic function**.

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logit($\hat{p}_d$) vs d



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Logistic regression

$$\ln \left(\frac{p_d}{1 - p_d} \right) = \beta_0 + \beta_1 d$$

Try least squares, regressing $\ln \left(\frac{\hat{p}_d}{1 - \hat{p}_d} \right)$ on the dose, d ?

Problems:

- What if $\hat{p}_d = 0$ or 1 ?
- $SD(\hat{p}_d)$ is not constant with d .

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Maximum likelihood

Assume $y_d \sim \text{Binomial}(n_d, p_d)$, y_d independent

with $\text{logit}(p_d) = \ln \left(\frac{p_d}{1 - p_d} \right) = \beta_0 + \beta_1 d$

Note:
$$p_d = \frac{e^{\beta_0 + \beta_1 d}}{1 + e^{\beta_0 + \beta_1 d}}$$

Likelihood:

$$L(\beta_0, \beta_1 | y) = \prod_d p_d^{y_d} (1 - p_d)^{(n_d - y_d)}$$

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Logistic regression in R

Logistic regression is a special case of a “generalized linear model”.

Function in R: `glm()`

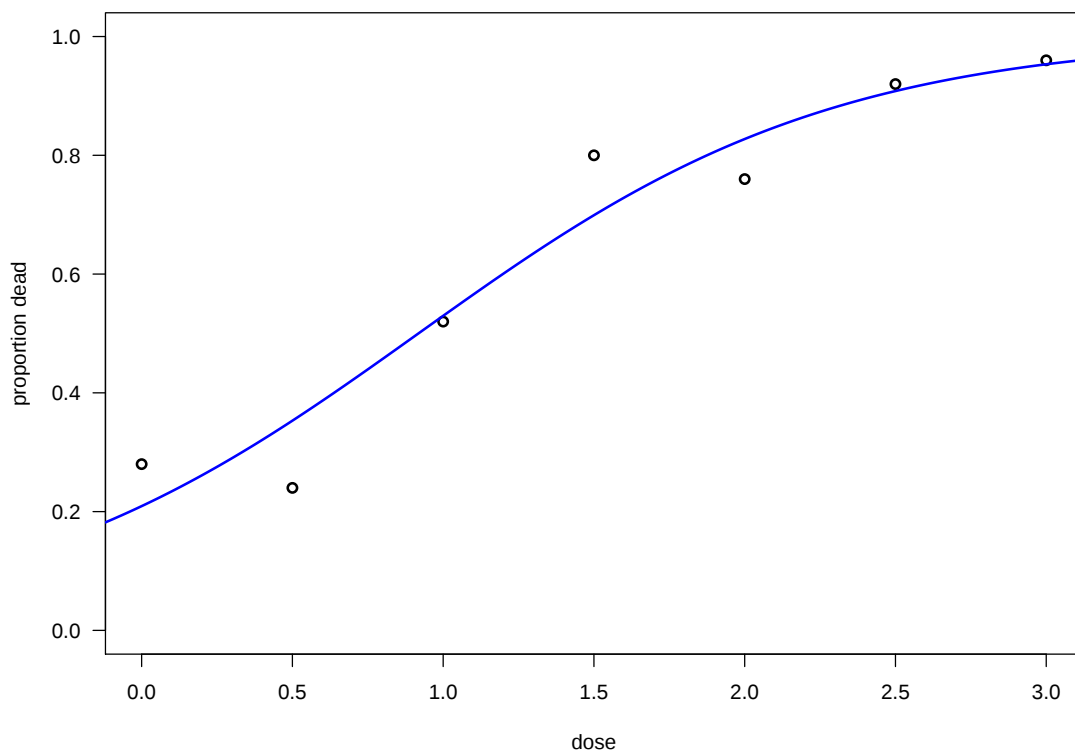
```
> glm.out <- glm(n.dead/n ~ dose, weights=n, data=spiders,  
                 family=binomial(link=logit))
```

```
> summary(glm.out)$coef
```

	Est	SE	t-val	P-val
(Intercept)	-1.33	0.33	-4.06	<0.001
dose	1.44	0.23	6.29	<0.001

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Fitted curve



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Interpretation of β 's

$$\ln \left(\frac{p_d}{1 - p_d} \right) = \beta_0 + \beta_1 d$$

β_0 = log odds when dose = 0

Note: $\beta_0 = 0 \longrightarrow p_0 = \frac{1}{2}$

β_1 = change in log odds with unit increase in dose

Note: $\beta_1 = 0 \longrightarrow$ survival unrelated to dose.

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LD50

LD50 = dose at which $\Pr(\text{dead} \mid \text{dose}) = \frac{1}{2}$.

$$\ln \left(\frac{1/2}{1 - 1/2} \right) = \beta_0 + \beta_1(\text{LD50})$$

$$\implies 0 = \beta_0 + \beta_1(\text{LD50})$$

$$\implies \text{LD50} = -\beta_0/\beta_1$$

$$\widehat{\text{LD50}} = -\hat{\beta}_0/\hat{\beta}_1$$

$$\widehat{\text{SE}}(\widehat{\text{LD50}}) \approx |\widehat{\text{LD50}}| \sqrt{\left(\frac{\widehat{\text{SE}}(\hat{\beta}_0)}{\hat{\beta}_0} \right)^2 + \left(\frac{\widehat{\text{SE}}(\hat{\beta}_1)}{\hat{\beta}_1} \right)^2 - 2 \frac{\text{cov}(\hat{\beta}_0, \hat{\beta}_1)}{\hat{\beta}_0 \hat{\beta}_1}}$$

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Estimating LD50 in R

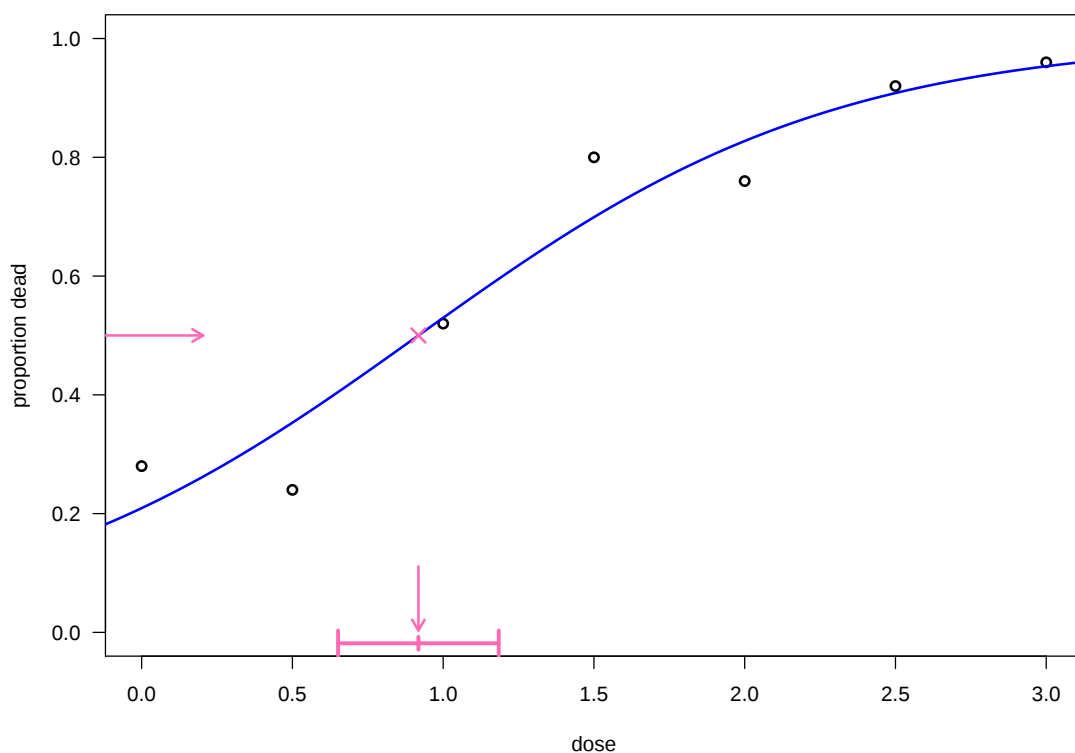
```
> glm.out <- glm(n.dead/n ~ dose, weights=n, data=spiders,
  family=binomial(link=logit))
> glm.sum <- summary(glm.out)

> co <- glm.out$coef
> ld50 <- -co[1]/co[2]
> se.co <- glm.sum$coef[,2]
> cov.co <- glm.sum$cov.scaled[1,2]
> se.ld50 <- abs(ld50) * sqrt( (se.co[1]/co[1])^2 +
  (se.co[2]/co[2])^2 -
  2*cov.co/(co[1]*co[2]) )

> ld50
  0.92
> se.ld50
  0.14
> ld50 + c(-1,1) * qnorm(0.975) * se.ld50
  0.65  1.18
```

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LD50



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Another example

Tobacco budworm, *Heliothis virescens*

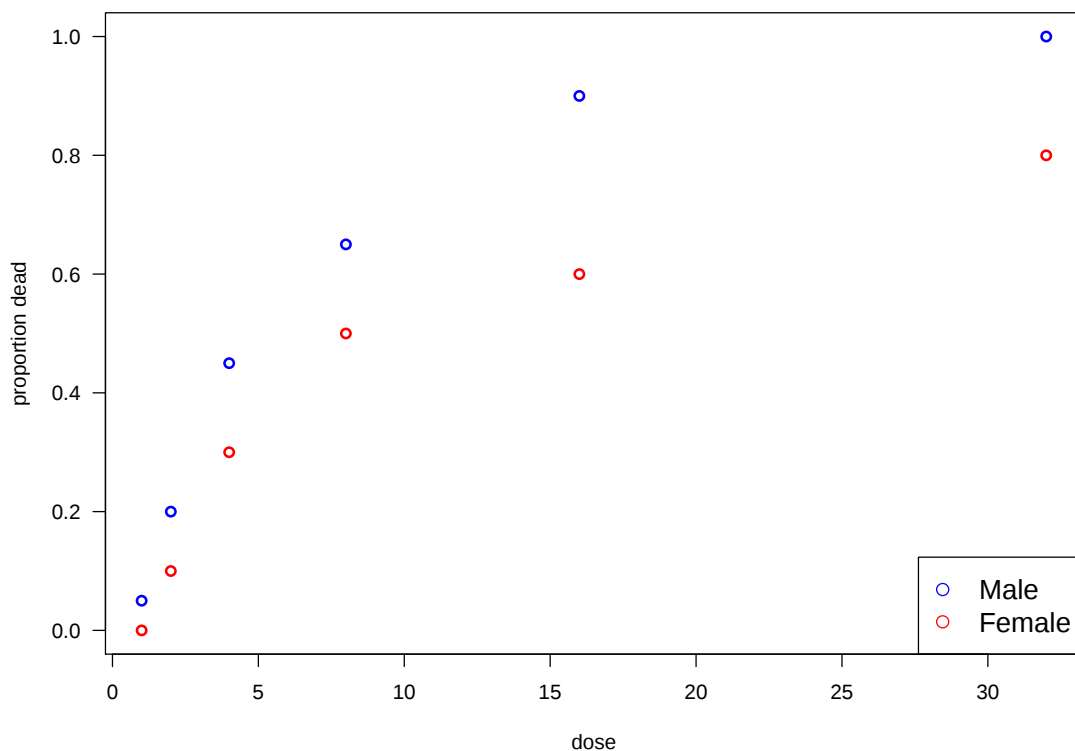
Batches of 20 male and 20 female worms were given a 3-day dose of pyrethroid *trans*-cypermethrin

The no. dead or “knocked down” in each batch was noted.

Sex	Dose					
	1	2	4	8	16	32
Male	1	4	9	13	18	20
Female	0	2	6	10	12	16

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A plot of the data



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Analysis in R

Assume no sex difference

```
> glm.out <- glm(n.dead/n ~ dose, weights=n, data=worms,  
                family=binomial(link=logit))
```

```
> summary(glm.out)$coef
```

	Est	SE	t-val	P-val
(Intercept)	-1.57	0.23	-6.8	<0.001
dose	0.153	0.022	6.8	<0.001

Assume sexes completely different

```
> glm.outB <- glm(n.dead/n ~ sex*dose, weights=n, data=worms,  
                  family=binomial(link=logit))
```

```
> summary(glm.outB)$coef
```

	Est	SE	t-val	P-val
(Intercept)	-1.72	0.32	-5.3	<0.001
sexmale	-0.21	0.51	-0.4	0.68
dose	0.116	0.024	4.9	<0.001
sexmale:dose	0.182	0.067	2.7	0.007

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Analysis in R (continued)

Different slopes but common “intercept”

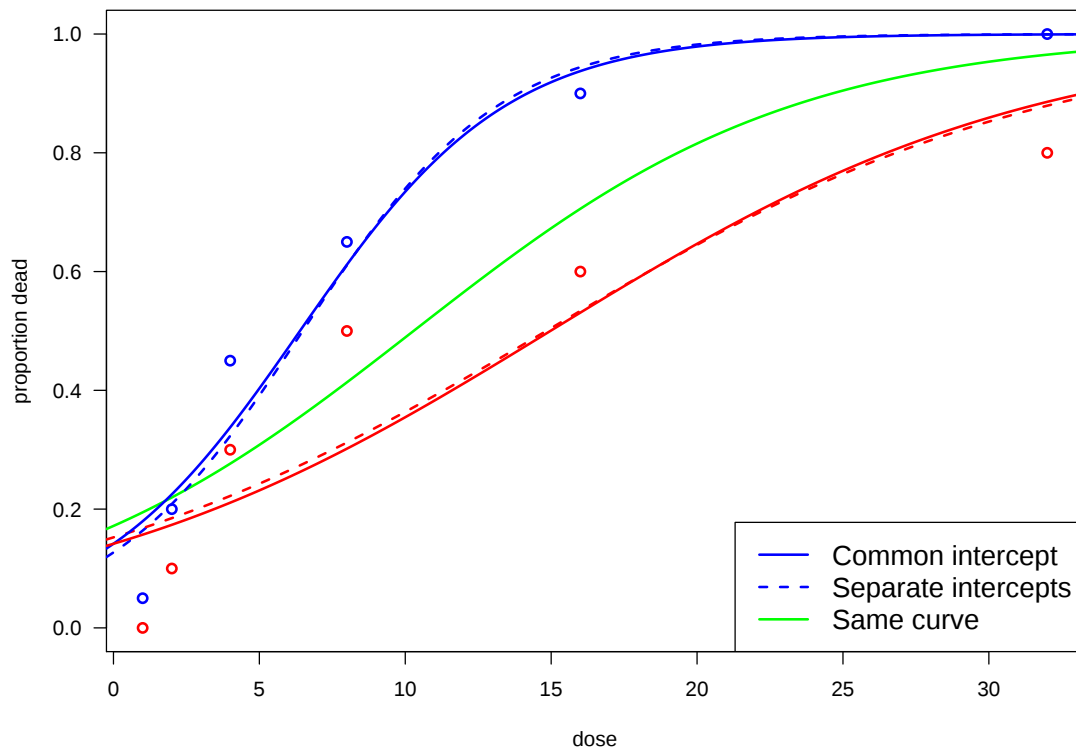
```
> glm.outC <- glm(n.dead/n ~ dose + sex:dose, weights=n,  
                  data=worms, family=binomial(link=logit))
```

```
> summary(glm.out)$coef
```

	Est	SE	t-val	P-val
(Intercept)	-1.80	0.25	-7.2	<0.001
dose	0.120	0.021	5.6	<0.001
dose:sexmale	0.161	0.044	3.7	<0.001

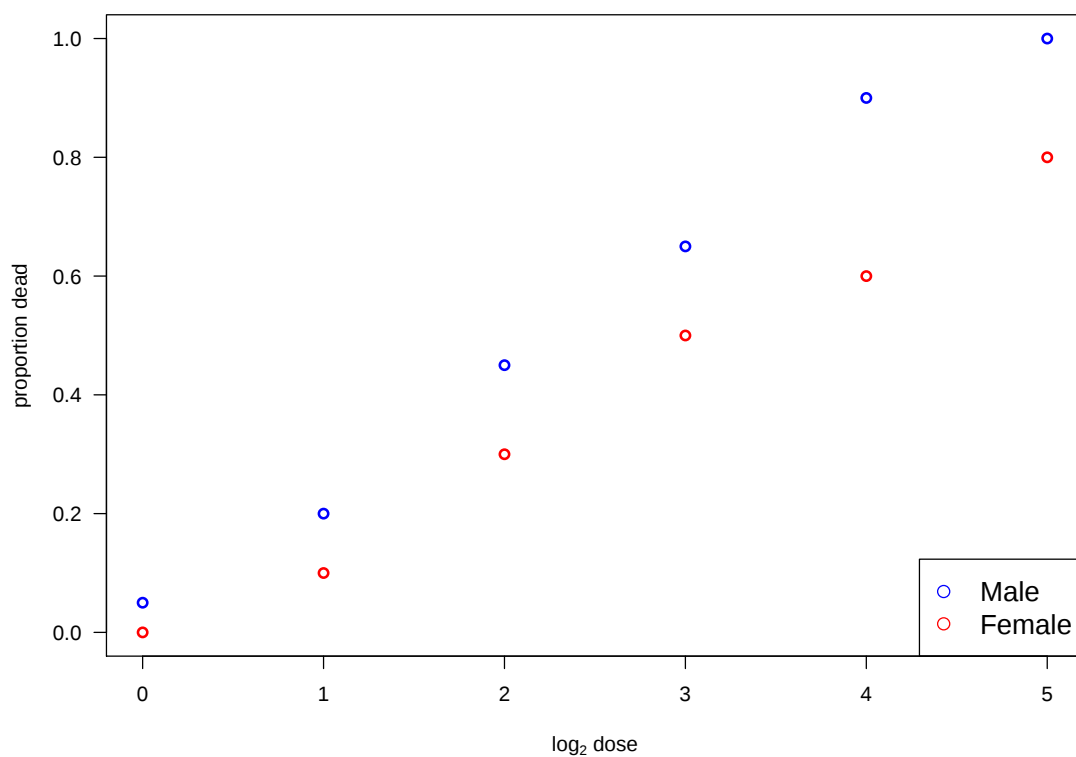
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Fitted curves



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Plot using $\log_2$ dose



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Use $\log_2$ of the dose

Assume no sex difference

```
> glm.out <- glm(n.dead/n ~ dose, weights=n, data=worms,
  family=binomial(link=logit))

> summary(glm.out)$coef
```

	Est	SE	t-val	P-val
(Intercept)	-2.77	0.37	-7.6	<0.001
dose	1.01	0.12	8.1	<0.001

Assume sexes completely different

```
> glm.outB <- glm(n.dead/n ~ sex*dose, weights=n, data=worms,
  family=binomial(link=logit))

> summary(glm.outB)$coef
```

	Est	SE	t-val	P-val
(Intercept)	-2.99	0.55	-5.4	<0.001
sexmale	0.17	0.78	-0.2	0.82
dose	0.91	0.17	5.4	<0.001
sexmale:dose	0.35	0.27	1.3	0.19

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Use $\log_2$ of the dose (continued)

Different slopes but common “intercept”

```
> glm.outC <- glm(n.dead/n ~ dose + sex:dose, weights=n,
  data=worms, family=binomial(link=logit))

> summary(glm.outC)$coef
```

	Est	SE	t-val	P-val
(Intercept)	-2.91	0.39	-7.5	<0.001
dose	0.88	0.13	6.9	<0.001
dose:sexmale	0.41	0.12	3.3	0.001

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Fitted curves

